

Chi-squared = "Badness of Fit"

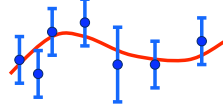
$$\chi^2 \equiv \sum_{i=1}^N \left(\frac{X_i - \mu_i(\alpha)}{\sigma_i} \right)^2 \sim \chi_{N-M}^2$$

X_i = data values $i = 1 \dots N$

σ_i = $1 - \sigma$ error bar

$\mu_i(\alpha)$ = model predicted data value

α_k = parameters of the model $k = 1 \dots M$



N = number of data points

M = number of fitted parameters

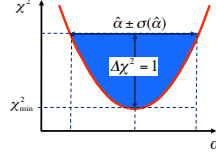
$N - M$ = degrees of freedom

The Dancing χ^2 Landscape

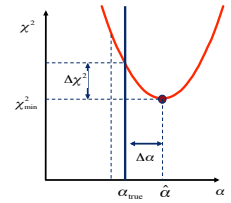
Fit M parameters to N data points.

$$\chi^2(X, \sigma, \alpha) = \sum_{i=1}^N \left(\frac{X_i - \mu_i(\alpha)}{\sigma_i} \right)^2$$

Best-fit parameters $\hat{\alpha}$ minimise χ^2 .



$$\sigma^2(\hat{\alpha}) = \frac{2}{\left(\frac{\partial \chi^2}{\partial \alpha} \right)_{\alpha=\hat{\alpha}}}$$



$$\hat{\alpha} \sim G(\alpha_{\text{true}}, \sigma^2(\hat{\alpha}))$$

$$\chi^2(\alpha_{\text{true}}) \sim \chi_N^2$$

$$\chi_{\text{min}}^2 \equiv \chi^2(\hat{\alpha}) \sim \chi_{N-M}^2$$

$$\Delta \chi^2 = \chi^2(\alpha_{\text{true}}) - \chi_{\text{min}}^2 \sim \chi_M^2$$

χ^2 distribution N degrees of freedom

$$f(x) = \frac{1}{\Gamma(N/2) 2^{N/2}} x^{(N/2-1)} e^{-x/2}$$

$$\Gamma(1) = 1 \quad \Gamma(1/2) = \sqrt{\pi}$$

$$\Gamma(n) = (n-1)! \quad \Gamma(\alpha+1) = \alpha \Gamma(\alpha)$$

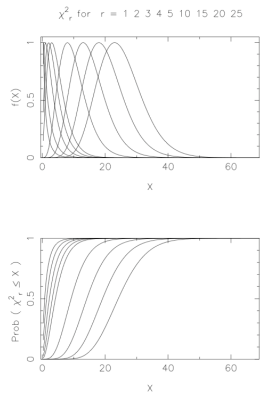
e.g. $\Gamma(3/2) = (1/2) \Gamma(1/2) = \sqrt{\pi}/2$

$$\chi_1^2: f(x) = \left(\frac{e^{-x}}{2\pi x} \right)^{1/2}$$

$$\chi_2^2: f(x) = \frac{1}{2} e^{-x/2}$$

$$\langle \chi_N^2 \rangle = N$$

$$\sigma^2(\chi_N^2) = 2N$$



Constructing χ_N^2 from N Gaussians

- Sum of squares of N independent Gaussian random variables

$\chi_N^2 \equiv$ Chi-squared with N degrees of freedom

X and Y are independent Gaussian random variables.

$$X \sim G(0,1) \quad Y \sim G(0,1)$$

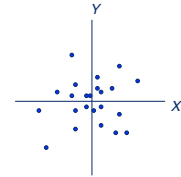
$$X^2 \sim \chi_1^2 \quad Y^2 \sim \chi_1^2$$

$$X^2 + Y^2 \sim \chi_2^2$$

and so on for each

new degree of freedom:

$$\chi_N^2 + \chi_M^2 \sim \chi_{N+M}^2$$



Data points with no error bars ☹



N data points: $\langle X_i \rangle = \langle X \rangle$ $\text{Cov}(X_i, X_j) = \sigma^2 \delta_{ij}$

Sample mean: $\bar{X} = \frac{1}{N} \sum X_i$ unbiased: $\langle \bar{X} \rangle = \langle X \rangle$

But σ_i unknown. How to estimate σ ?

Variance: $\sigma^2(X) = \langle (X - \langle X \rangle)^2 \rangle$

$$\text{Try: } s^2 = \frac{1}{N} \sum (X_i - \bar{X})^2$$

Is $\langle s^2 \rangle = \sigma^2$?

No. $\langle s^2 \rangle < \sigma^2$ Must correct for a bias.

Sample Variance S^2 : Unbiased for σ^2



$$S^2 = A \sum_{i=1}^N (X_i - \bar{X})^2 \quad \text{Pick } A \text{ so that } \langle S^2 \rangle = A \sum_{i=1}^N \langle (X_i - \bar{X})^2 \rangle = \sigma^2$$

$$\langle (X_i - \bar{X})^2 \rangle = \langle [(X_i - \langle X \rangle) - (\bar{X} - \langle X \rangle)]^2 \rangle$$

$$= \langle (X_i - \langle X \rangle)^2 - 2(X_i - \langle X \rangle)(\bar{X} - \langle X \rangle) + (\bar{X} - \langle X \rangle)^2 \rangle$$

$$= \sigma^2(X_i) - 2 \text{Cov}(X_i, \bar{X}) + \sigma^2(\bar{X})$$

$$= \sigma^2 - 2 \frac{\sigma^2}{N} + \frac{\sigma^2}{N}$$

Note: $\text{Cov}(X_i, \bar{X}) = \frac{\sigma^2}{N}$

$$= \left(1 - \frac{1}{N} \right) \sigma^2 = \left(\frac{N-1}{N} \right) \sigma^2$$

$$\therefore \langle S^2 \rangle = A(N-1) \sigma^2 \quad \text{Pick } A = \frac{1}{N-1}$$

$$S^2 \equiv \frac{1}{N-1} \sum_{i=1}^N (X_i - \bar{X})^2$$

Evaluation of Cov(X_i, \bar{X})

$$\text{Cov}(X_i, \bar{X}) = \langle (X_i - \langle X_i \rangle) (\bar{X} - \langle \bar{X} \rangle) \rangle$$

Note : $\langle X_i \rangle = \langle \bar{X} \rangle = \langle X \rangle$

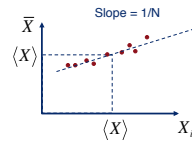
Shift coords to put $\langle X \rangle = 0$:

$$\text{Cov}(X_i, \bar{X}) = \langle (X_i - 0) (\bar{X} - 0) \rangle$$

$$= \left\langle X_i \frac{1}{N} \sum_k X_k \right\rangle$$

$$= \frac{1}{N} \sum_k \langle X_i X_k \rangle$$

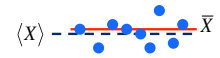
$$= \frac{1}{N} \sum_k \sigma^2 \delta_{ik} = \frac{\sigma^2}{N}$$



$$\text{Cov}(X_i, X_j) = \sigma^2 \delta_{ij}$$

Sample Variance S^2 : Unbiased for σ^2

$$S^2 = \frac{1}{N-1} \sum_{i=1}^N (X_i - \bar{X})^2$$



Why $\frac{1}{N-1}$, not $\frac{1}{N}$?

Because \bar{X} "chases" the dancing data points, removing 1 "degree-of-freedom" from the dance.

$$S^2 \sim \frac{\sigma^2}{N-1} \chi_{N-1}^2$$

$$\langle S^2 \rangle = \frac{\sigma^2}{N-1} \langle \chi_{N-1}^2 \rangle$$

$$= \frac{\sigma^2}{N-1} (N-1) = \sigma^2$$

$$\text{Var}[S^2] = \left(\frac{\sigma^2}{N-1} \right)^2 \text{Var}[\chi_{N-1}^2]$$

$$= \left(\frac{\sigma^2}{N-1} \right)^2 2(N-1) = \frac{2\sigma^4}{N-1}$$

$$\frac{\sigma(S^2)}{S^2} = \left(\frac{2}{N-1} \right)^{1/2} = \text{fractional accuracy}$$

Degrees of Freedom (DoF)

N data points: $\langle X_i \rangle = \langle X \rangle$ $\text{Cov}(X_i, X_j) = \sigma^2 \delta_{ij}$

$$\sum_{i=1}^N \left(\frac{X_i - \langle X \rangle}{\sigma} \right)^2 \sim \chi_N^2 \quad N \text{ degrees of freedom.}$$

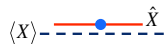


If $\langle X \rangle$ unknown, use \hat{X} instead:

$$\sum_{i=1}^N \left(\frac{X_i - \hat{X}}{\sigma} \right)^2 \sim \chi_{N-1}^2 \quad N-1 \text{ degrees of freedom.}$$

If $N=1$ data point: $\hat{X}=X_1$

$$\left(\frac{X_1 - \langle X \rangle}{\sigma} \right)^2 \sim \chi_1^2 \quad 1 \text{ degree of freedom}$$



$$\left(\frac{X_1 - \hat{X}}{\sigma} \right)^2 = 0 \quad 0 \text{ degrees of freedom.}$$

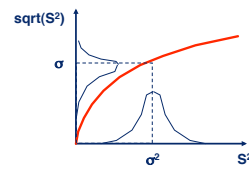
Fit M parameters to N data points:

$$\sum_{i=1}^N \left(\frac{X_i - \mu(\alpha)}{\sigma} \right)^2 \sim \chi_{N-M}^2 \quad N-M \text{ degrees of freedom.}$$

Each fitted parameter removes 1 degree of freedom from the residuals:

sqrt(S^2) is biased

- The sample variance S^2 is unbiased for σ^2 .
- Is the sqrt(S^2) unbiased for σ ?
- No. The square root introduces a bias:



$$\langle \sqrt{S^2} \rangle < \sigma, \text{ even though } \sqrt{\sigma^2} = \sigma.$$

Example: Correct the Bias in sqrt(S^2)

Define $g(x) = \text{sqrt}(x)$ $\bar{S} = g(S^2)$

and its derivatives:

$$g(x) = x^{1/2}, \quad g'(x) = \frac{1}{2} x^{-1/2}, \quad g''(x) = -\frac{1}{4} x^{-3/2}$$

Evaluate the bias:

$$\langle \bar{S} \rangle = g(\langle S^2 \rangle) + \frac{g''(\langle S^2 \rangle)}{2} \text{Var}(S^2) + \dots$$

$$= g(\sigma^2) + \frac{g''(\sigma^2)}{2} \frac{2\sigma^4}{N-1} + \dots$$

$$= \sigma - \frac{1}{8\sigma^3} \frac{2\sigma^4}{N-1} + \dots = \sigma \left(1 - \frac{1}{4(N-1)} + \dots \right) = \sigma \left(\frac{4N-5}{4N-4} \right) + \dots$$

$$\text{Bias - corrected: } \bar{S} = \left(\frac{4N-4}{4N-5} \right) (S^2)^{1/2}$$

"Robust" estimation methods

- Robust => less sensitive to "bad" data.

- Example: using **median** rather than **mean**:



- Sample Mean \bar{X} minimizes the Sample Variance:

- Median X_M minimizes the "Mean Absolute Deviation" :

$$S^2 = \frac{1}{N-1} \sum_{i=1}^N (X_i - \mu)^2$$

$$\text{MAD} = \frac{1}{N} \sum_{i=1}^N |X_i - \mu|$$

$$\frac{\partial}{\partial \mu} \left[\sum_{i=1}^N (X_i - \mu)^2 \right] = 0$$

$$\frac{\partial}{\partial \mu} \left[\sum_{i=1}^N |X_i - \mu| \right] = 0$$

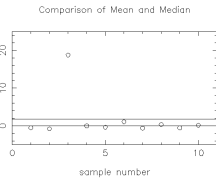
for $\mu = \bar{X}$

for $\mu = X_M = \text{Median}(X_i)$

Mean vs Median

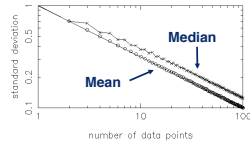
- The median is less sensitive to outliers than the mean.

Mean
Median



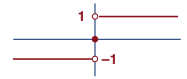
- The median is unbiased, but not a minimum-variance estimator.

- Note how the standard deviations of the median and of the mean vary with sample size.



“Proof” that the Median minimises the MAD

$$H(x) = \begin{cases} +1, & x > 0 \\ 0, & x = 0 \\ -1, & x < 0 \end{cases} \quad \frac{dH}{dx} = 2\delta(x)$$



$$MAD = \frac{1}{N} \sum_{i=1}^N |\mu - X_i| = \frac{1}{N} \sum_{i=1}^N (\mu - X_i) H(\mu - X_i)$$

$$\begin{aligned} \frac{dMAD}{d\mu} &= \frac{1}{N} \sum_{i=1}^N [H(\mu - X_i) + (\mu - X_i) H'(\mu - X_i)] \\ &= \frac{1}{N} \sum_{i=1}^N H(\mu - X_i) \quad \text{since } H'(x) = 0 \text{ whenever } x \neq 0 \\ &= 0 \quad \text{if } \mu = \text{median}(X_i) \end{aligned}$$

Finding the Median without Sorting

$$\frac{X_i - X_M}{|X_i - X_M|} = \begin{cases} 1 & \text{if } X_i > X_M \\ 0 & \text{if } X_i = X_M \\ -1 & \text{if } X_i < X_M \end{cases}$$

- A useful application:

Since $\sum_{i=1}^N \frac{X_i - X_M}{|X_i - X_M|} = 0$, first make a guess at X_M .

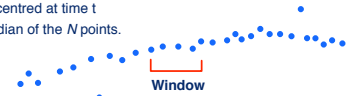
$$\text{Then estimate a new } X_M = \frac{\sum_{i=1}^N \frac{X_i}{|X_i - X_M|}}{\sum_{i=1}^N \frac{1}{|X_i - X_M|}}$$

and iterate to convergence.

Median Filter and Sigma-Clip

Median filter:

- window of N points centred at time t
- medfilt(t) is the median of the N points.



Sigma-clip:

- Fit all points by minimising χ^2
- Set threshold K and check for outliers at $\pm K\sigma$ or more
- Repeat fit omitting largest outlier
- Iterate until set of rejected points converges.



Various “Badness-of-Fit” Statistics

Statistic	Optimal Value	Badness Function	Sigma-clip
Sample Variance $S^2 = \frac{1}{N-1} \sum_{i=1}^N (X_i - \mu_i)^2$	mean \bar{X}	ϵ^2	Sigma-clip $\pm K\sigma$
Chi-squared $\chi^2 = \sum_{i=1}^N \left(\frac{X_i - \mu_i}{\sigma_i} \right)^2$	optimal average \hat{X}	η^2	
Mean Absolute Deviation $MAD = \frac{1}{N} \sum_{i=1}^N X_i - \mu_i $	median X_M	$ \epsilon $	
Sum Absolute Normalised Errors $SANE = \sum_{i=1}^N \left \frac{X_i - \mu_i}{\sigma_i} \right $		$ \eta $	

